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The triple-flash illusion reveals a driving role of alpha-band reverberations in visual perception

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The triple-flash illusion reveals a driving role of alpha-band reverberations in visual perception

RUNNING TITLE: Driving role of alpha oscillations in perception

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32 ABSTRACT

33 The modulatory role of spontaneous brain oscillations on perception of threshold-level stimuli is well
34 established. Here, we provide evidence that alpha-band (~10 Hz) oscillations not only modulate
35 perception of threshold-level sensory inputs, but also can drive perception and generate percepts without a
36 physical stimulus being present. We used the “triple-flash” illusion: Occasional perception of three
37 flashes when only two spatially-coincident veridical ones, separated by ~100 ms, are presented. The
38 illusion was proposed to result from superposition of two hypothetical oscillatory impulse response
39 functions (IRF) generated in response to each flash: When the delay between flashes matches the period
40 of the oscillation, the superposition enhances a later part of the oscillation that is normally damped; when
41 this enhancement crosses perceptual threshold, a third flash is erroneously perceived (Bowen, 1989). In
42 Experiment 1, we varied stimulus onset asynchrony (SOA) and validated Bowen’s theory: the optimal
43 SOA for illusion to occur was correlated, across human subjects (both genders), with the subject-specific
44 IRF period determined from a separate EEG experiment. Experiment 2 revealed that pre-stimulus parietal
45 – but no occipital – alpha EEG phase and power, as well as post-stimulus alpha phase-locking, together
46 determine the occurrence of the illusion on a trial-by-trial basis. Thus, oscillatory reverberations create
47 something out of nothing – a third flash where there are only two.

48 SIGNIFICANCE

49 We highlight a novel property of alpha-band (~10 Hz) oscillations based on three experiments (two EEG
50 and one psychophysics) by demonstrating that alpha-band oscillations do not merely modulate perception,
51 but can also drive perception. We show that human participants report seeing a third flash when only two
52 are presented (the “triple-flash” illusion) most often when the inter-flash delay matches the period of
53 participant’s oscillatory impulse response function reverberating in alpha. Within-subject, the phase and
54 power of ongoing parietal—but not occipital—alpha-band oscillations at the time of the first flash
55 determine illusory percept on a trial-by-trial basis. We revealed a physiologically plausible mechanism
56 that validates and extends the original theoretical account of the triple-flash illusion proposed by Bowen
57 in 1989.

58 **INTRODUCTION**

59 Spontaneous rhythmic fluctuations in various frequency bands have been consistently reported to
60 affect perception across several sensory modalities (VanRullen, 2016b). Most studies found modulatory
61 effects of pre-stimulus alpha-band (~10 Hz) power, phase, and frequency on threshold-level stimulus
62 detection, perceptual and temporal discrimination (Busch et al., 2009; Roberts et al., 2014; Baumgarten et
63 al., 2015; Samaha and Postle, 2015). These perceptual consequences of endogenous oscillations imply
64 that perception is inherently rhythmic and operates in a form of perceptual cycles, with periods of high
65 and low excitability (Dugue et al., 2011; VanRullen et al., 2011). However, most evidence for rhythmicity
66 in perception is based on modulatory effects of ongoing oscillations on stimulus processing: For example,
67 oscillatory phase at stimulus onset modulates stimulus visibility by approximately 10-15% (Mathewson et
68 al., 2009; Busch and VanRullen, 2010; Dugue et al., 2011). Do brain oscillations merely modulate
69 perception, or can they also, under certain conditions, drive perception over perceptual threshold and
70 generate a percept without a physical stimulus?

71 In the “triple-flash” visual illusion, two brief light pulses separated by about 100 ms are
72 sometimes perceived as three (Bowen, 1989). The effect was theoretically explained by the superposition
73 of two damped oscillatory impulse response functions (IRF) generated in response to each stimulus flash:
74 The third illusory flash is perceived when the delay between veridical flashes matches the period of
75 oscillatory IRF. In this case, the later part of the oscillation is enhanced, and when this enhancement
76 crosses perceptual threshold a third illusory flash is perceived (Fig. 1B, middle panel). However, when
77 the delay does not match the period, the later part of the oscillation is dampened, and only two flashes are
78 perceived (Fig. 1B, top and bottom panels). Thus, Bowen’s theoretical account of the triple-flash illusion
79 assumes that the brain response to a single flash of light is oscillatory.

80 Empirical tests of Bowen’s insightful predictions, and the role of oscillations in the triple-flash
81 illusion, have not yet been demonstrated. Several findings, however, indicate that alpha-band oscillations
82 could be involved in the generation of illusory third-flash percepts. First, the optimal two-flash delay of

~100 ms corresponds to the average period of alpha-band oscillations. Additionally, the optimal inter-flash delay for the illusion to be perceived varies across individuals (Bowen, 1989), and so does the alpha peak frequency across individuals (IAF) and across brain areas within individual (Doppelmayr et al., 1998; Haegens et al., 2014). Second, variations in occipital IAF are causally linked to the temporal properties of visual perception, such that faster occipital alpha oscillations are associated with finer temporal resolution in perception (Cecere et al., 2015; Samaha and Postle, 2015). Third, there is empirical evidence that the response to a single flash is indeed oscillatory, and reverberates at ~10 Hz (VanRullen and Macdonald, 2012).

The triple-flash illusion appears similar to other phantom-flash illusions (Apthorp et al., 2013), e.g. the sound-induced double-flash illusion, whose temporal window is causally related to alpha-band oscillations (Cecere et al., 2015). However, the “triple-flash” illusion is purely endogenous, whereas perception of other phantom-flash illusions requires simultaneous presentation of additional stimuli, either in a different modality or in a different spatial location.

Understanding the role of oscillations in the triple-flash illusion is critical to the idea of perceptual cycles, as the illusory third-flash is potentially caused by perceptual reverberations. According to this line of reasoning, cortical excitability and stimulus-evoked responses determined by the power and phase of ongoing alpha-band oscillations at the moment of the first flash could have carryover effects for several alpha cycles (Remond and Lesevre, 1967; Jansen and Brandt, 1991; Busch et al., 2009; Mathewson et al., 2009; Fiebelkorn et al., 2013). In Experiment 1, we directly tested whether illusory third-flash perception could result from summed reverberations of visual responses as proposed by Bowen (1989). In Experiment 2, we investigated pre-stimulus and stimulus-related effects of alpha phase, power, and frequency (separately for occipital and parietal alpha sources) on the illusory third-flash perception.

109 **METHODS**110 **EXPERIMENT 1**

111 This experiment consisted of two parts. In the first part, we estimated subject-specific optimal SOA of the
 112 triple-flash illusion using a psychophysical approach. In the second part, we established the subject-
 113 specific IRF using EEG recordings obtained during white noise flicker stimulation.

114 *Participants.* Thirty participants (12 females, mean age 27.4) took part in the psychophysics experiment.
 115 All participants had normal or corrected-to-normal vision. The data from three participants was excluded
 116 from the analyses: Two due to low accuracy on easy-to-discriminate long SOA three-flash trials, and one
 117 due to bias towards reporting illusions on very short SOA trials, with no discernable preferred SOA for
 118 the triple-flash illusion to be perceived. The study was conducted in accordance with the declaration of
 119 Helsinki and approved by the local ethics committee. Written informed consent was obtained from all
 120 participants prior to the experiment.

121 *Stimuli and Procedure.* Stimuli for the triple-flash experiment were two or three uniform white circles
 122 (radius 3.5°) presented on a black background in rapid succession peripherally above the fixation dot
 123 (eccentricity 7° visual angle). The viewing distance was constrained by the chin rest positioned 57 cm
 124 away from a 17-in. CRT monitor (600 x 480 pixels; 160 Hz vertical refresh rate). On each trial, either two
 125 or three circles were presented for a duration of one screen refresh (6.25 ms) with variable stimulus onset
 126 asynchrony (SOA; Fig. 1A). After the offset of the last flash, the screen remained black until a response
 127 was made. Following the response, the next trial started after a variable delay (1000–1500 ms). For a
 128 majority of participants ($N = 20$), the SOA on two-flash trials was randomly selected from one of 10
 129 possible SOAs (50, 62.5, 75, 87.5, 100, 112.5, 125, 137.5, 150, 175 ms), with 80 trials per SOA; and on
 130 three-flash trials from 8 possible SOAs (25, 31.25, 37.5, 50, 75, 87.5, 175, and 250 ms), with 50 trials per
 131 SOA. For the remaining 10 participants, a finer sampling of SOAs was used: For the two-flash trials – 17
 132 SOAs (50 to 250 ms, in steps of 12.5 ms), with 40 trials per SOA; and for the three-flash trials – 27 SOAs
 133 (25 to 125 ms, in steps of 6.25 ms, and 125 to 250 ms, in steps of 12.5 ms), with 20 trials per SOA. The

134 overall probability of three flash trials was 33% (coarsely sampled SOAs) or 44% (finely sampled SOAs)
 135 throughout the experiment. The task consisted of 20 practice trials and 1200 (coarsely sampled SOAs) or
 136 1220 (finely sampled SOAs) experimental trials separated in blocks of 100 trials. Participants were tested
 137 in a dark room. They were instructed to keep their eyes on the fixation dot throughout the experiment, and
 138 to make speeded responses without sacrificing accuracy. “Left arrow” and “right arrow” keys were used
 139 to indicate perception of two- and three-flashes respectively. Participants were encouraged to take breaks
 140 after each block. The experiment lasted approximately an hour.

141 In a separate experimental session, we recorded EEG while participants monitored a peripheral
 142 disk stimulus, whose luminance randomly varied every screen refresh with the constraint that the power
 143 spectrum of the luminance time course was flat between 0 and 80 Hz (i.e., white noise). Stimulus size,
 144 position, viewing distance, background, and screen refresh rate were kept identical to the triple-flash
 145 experimental session. Participants were instructed to detect target stimuli (a small circle surrounded by a
 146 darker annulus) presented in the center of the flickering stimulus by pressing a button. There were 2-4
 147 targets presented during each 6.25 sec long trial. Target stimulus detectability (set to 50% using adaptive
 148 staircase procedure based on the performance during the first 30 trials) was manipulated by changing the
 149 relative contrast between the small circle and the annulus. The beginning of each trial was self-paced
 150 using a button press. The experiment lasted about one hour, and was divided in 8 blocks, with 48 trials in
 151 each block. Participants were encouraged to take rest breaks after each block. Further details on the
 152 rationale of this EEG study, and additional analyses can be found in Brüers and VanRullen (2017).

153 *Data analysis.* To evaluate whether the triple-flash perception is related to oscillatory IRF, we correlated
 154 the period of subject-specific IRF with the subject-specific optimal SOA for illusory perception.
 155 Robustness of correlations was assessed by performing a bootstrapping procedure (resampling with
 156 replacement) using the Robust Correlation Toolbox (Pernet et al., 2012).

157 The period of subject-specific IRF was obtained by carrying out the following analysis steps.
 158 First, EEG data were pre-processed using the same pipeline as for the data in the Experiment 2 (see

below), with a few exceptions: (1) data were downsampled to 160 Hz to match the rate of stimulus luminance changes, (2) epochs were -250 – 6500 ms relative to the onset of the luminance sequence, (3) trials containing eye blink artefacts during luminance sequences were excluded from the analyses. Second, single-trial EEG time-series were cross-correlated with stimulus luminance time series at lags between -400 and 1300 ms in steps of 1/160. Third, the FFT of the trial-average cross-correlation result in the -400 - 1300 ms was performed. The data was zero-padded to obtain 0.1 Hz frequency resolution. The period of subject-specific IRF was determined by finding the peak frequency (f) in the range of 6-14 Hz, and expressing it as period in milliseconds ($1/f$).

Subject-specific optimal SOA for the triple-flash illusion to be perceived was determined by fitting symmetrical and non-symmetrical functions to the behavioral performance on two-flash trials (proportion of two-flash trials perceived as three). Initial to-be-fitted model parameters were based on each subject's empirical data, and used as an input for *fminsearch* Matlab function. We used Gaussian (1), Weibull (2), and ex-Gaussian (3) functions:

$$(1) f(x) = ae^{-\frac{(x-b)^2}{2c^2}},$$

where a is Gaussian peak y-axis value (initially set to maximum number of illusions perceived by participant); b Gaussian peak x-axis value (initially set to SOA of maximum number of illusions); c is width of the Gaussian (initially set to 0.05 sec).

$$(2) f(x) = c \left(\frac{x}{a}\right)^{b-1} e^{-\left(\frac{x}{a}\right)^b},$$

where a is the x-axis scaling factor (initially set to SOA of maximum number of illusions perceived by participant); b defines the shape of the curve (initially set to 4); and c scales the curve along the y-axis (initially set to twice the maximum number of illusions).

$$(3) f(x) = h \frac{\lambda}{2} e^{\frac{\lambda}{2}(2\mu + \lambda\sigma^2 - 2x)} \operatorname{erfc}\left(\frac{\mu + \lambda\sigma^2 - x}{\sqrt{2}\sigma}\right)$$

where $\lambda = 1/\tau$, and τ is an exponential decay parameter (initially set to 0.1) ; μ is mean of the Gaussian (initially set to half the SOA with maximum number of illusions); σ is variance of the Gaussian (initially

183 set to 0.05 sec); h is the y-axis scaling factor (initially set to maximum number of illusions divided by
184 four).

185 Goodness of fit at each *fminsearch* iteration and across the three different fitting functions was
186 evaluated using R^2 , which is the amount of variance accounted for.

187

188 **EXPERIMENT 2**

189 The purpose of Experiment 2 was to characterize EEG changes that accompany perception of the third-
190 flash illusion relative to no-illusion trials with physically identical stimuli.

191 *Participants.* EEG data were recorded from 35 participants (18 females, mean age 26.5), 18 of them also
192 participated in Experiment 1. All participants had normal or corrected-to-normal vision. The study was
193 conducted in accordance with the declaration of Helsinki and approved by the local ethics committee.
194 Written informed consent was obtained from all participants prior to the experiment.

195 *Stimuli and Procedure.* In comparison to the procedure of Experiment 1, two-flash trial SOA was fixed at
196 87.5 ms. The choice of this SOA was based on the results of Experiment 1, where maximal number of
197 illusions on average across participants based on ex-Gaussian fits was 92 ms ($SD = 17$ ms). As in
198 Experiment 1, three-flash trial SOA was variable (25, 31.25, 37.5, 50, 75, 87.5, 175, and 250 ms), and
199 overall proportion of three-flash trials was kept at 33% throughout the experiment. To avoid muscle
200 artifacts in EEG, viewing distance was unconstrained by the chin rest but still kept at approximately 57
201 cm.

202 *Data acquisition and preprocessing.* EEG data were recorded using 64-channel ActiveTwo BioSemi
203 system (for detailed description of the hardware see www.biosemi.com) at 1024 Hz sampling rate.
204 Offline, the data were down-sampled to 512 Hz and re-referenced to the average activity over all
205 electrodes. Continuous EEG recordings were band-pass filtered at 0.5 – 80 Hz, and electrical line noise
206 was removed using a notch filter (band-stop 47 to 53 Hz). The data were then epoched (-1500 ms to 2000

ms relative to the first stimulus onset), and baseline-corrected with respect to the time window of $-200-0$ ms relative to the first stimulus onset.

Artifact removal was done in two steps. First, the data was visually inspected and trials containing muscle artifacts or eye blinks during and 800 ms prior to the stimulus presentation were removed. The second artifact rejection step involved independent components analysis (ICA; Delorme and Makeig, 2004). Components that did not account for any brain activity, such as eye-movements or noise, were subtracted from the data (on average, 1.4 components per subject). Furthermore, trials with reaction time (RT) of 3 standard deviations longer than subject's mean RT were excluded from the analysis. On average, 90.6 % of two-flash trials per subject were included in the analysis (SD = 4.4%).

Alpha-band source separation. To account for variability in alpha peak frequency across individuals, and across brain regions within an individual (Haegens et al., 2014), we separated parietal and occipital alpha sources based on ICA using JADE algorithm as implemented in EEGLAB and dipole fitting using DIPFIT plug-in of the EEGLAB toolbox (Oostenveld et al., 2003; Delorme and Makeig, 2004). First, eye-blink and movement artifact-free data was band-pass filtered at 5-15 Hz. The filter kernel was created using Matlab *firls* function with a filter order of 307 ($3 \times (\text{sampling rate} / \text{lower bound of the filter})$), and 15% transition zones (Cohen, 2014a). Thereafter, the ICA was performed on temporally-filtered data using only the pre-stimulus time window (-1000 to 0 ms, where 0 is the first stimulus onset), obtaining 20 independent components (ICs). For each independent component (IC), a single-equivalent current dipole model was fitted using three-layer BEM template model based on the standard Montreal Neurological Institute's (MNI) brain template from the DIPFIT plug-in. One parietal and one occipital IC was selected based on the spatial proximity of the fitted dipoles to the reference anatomical coordinates, with constraints that selected equivalent dipoles had less than 15% residual variance from the spherical forward-model scalp projection, and were located inside the model brain volume. Reference anatomical coordinates for parietal ROIs were centered on the left and on the right Brodmann area 7 (right-side MNI coordinates: $-20 -90 0$; left-side MNI coordinates: $20 -90 0$); for occipital ROIs, mean coordinates of

232 Brodmann areas 17 and 18 were used (right-side MNI coordinates: -20 -70 50; left-side MNI coordinates:
233 20 -70 50; Haegens et al., 2014).

234 Individual alpha-peak frequency (IAF) at occipital and parietal alpha sources was estimated by
235 taking the FFT of single-trial IC time series in the -1000 – 0 ms window. The data was zero-padded to
236 obtain 0.1 Hz frequency resolution. The absolute value of FFT coefficients was squared and averaged
237 across trials. The individual alpha-peak frequency was determined as the peak in the range of 6-14 Hz.
238 This frequency window was chosen based on a large-sample study (N=96) by Bazanova & Vernon
239 (2014), which demonstrated that individual alpha peak can vary from 6 to 14 Hz. To make sure that IAF
240 estimation using all trials was independent from the magnitude of alpha-band power, we conducted a
241 control analysis to test for differences in IAF on high vs. low alpha-band power trials. High- and low-
242 alpha power trials were defined by means of a median split of all trials with respect to alpha power (mean
243 in the 10±3 Hz frequency window) during pre-stimulus interval. Out of the total of 33 participants, on low
244 alpha power trials the parietal IAF could be determined for 30, and the occipital IAF for 28 participants
245 (IAF was determined using the same procedure as described above). There were no statistically
246 significant differences for IAF on high- vs. low alpha-band power trials neither for parietal ($t(29) = .136$,
247 $p = .892$), nor occipital alpha sources ($t(27) = .431$, $p = .670$). Given this result, and the fact that IAF
248 determined from all trials has higher signal-to-noise ratio and thus provides a more reliable estimate of
249 IAF, for between-subject analyses we used IAF defined from all trials.

250 *Independent component time-frequency analyses.* Epoched unfiltered data was multiplied by ICA
251 unmixing matrix from selected subject-specific parietal and occipital components to obtain component
252 single-trial time series. Time-frequency decomposition was performed by convolving single-trial data
253 from parietal and occipital ICs with complex Morlet wavelets, defined as:

$$e^{i2\pi f_i t} e^{-t^2/(2\sigma^2)}$$

254 where t is time, f_i is frequency which ranged from 2 to 40 Hz in 39 logarithmically spaced steps, and σ is
255 the width of each frequency band defined as $n/(2\pi f_i)$, where n is a number of wavelet cycles that varied

from 1 to 7 in logarithmically spaced steps to obtain comparable frequency precision at low and high frequencies.

Instantaneous power was computed as the square of the analytic signal Z (a complex result of convolution) and averaged across trials (i.e., $power = \text{Re}[Z(t)]^2 + \text{Im}[Z(t)]^2$). Thus obtained power values were then baseline-corrected by converting to decibel scale ($10 \log_{10}(power/baseline)$), where condition-average power from -400 to -100 ms pre-stimulus period served as the baseline. Condition-average rather than condition-specific baseline was used to avoid introducing spurious power differences in the post-stimulus period.

To evaluate the effects of alpha phase on illusory third-flash perception, we compared phase distributions on illusion vs. non-illusion two-flash trials using phase opposition sum (POS), which is expressed as a function of inter-trial phase clustering (ITPC) of illusion and non-illusion trials relative to ITPC of all trials (VanRullen, 2016a):

$$POS = ITPC_{\text{illusion_trials}} + ITPC_{\text{non-illusion trials}} - 2 ITPC_{\text{All trials}}$$

where $ITPC = \left| \frac{1}{n} \sum_{j=1}^n e^{i\varphi_t} \right|$, the phase angle at each time point $\varphi_t = \arctan(\text{Im}[Z(t)]/\text{Re}[Z(t)])$, n is the number of trials, j is the trial, and i is the complex operator.

To test the hypothesis that third-flash perception is related to more precise phase alignment of oscillatory responses to veridical flashes, we computed phase consistency across trials at 11.43 Hz (corresponding to the 87.5 ms SOA) using the weighted pair-wise phase consistency (wPPC) metric (Vinck et al., 2010). Consistency of phases across trials is often estimated using ITPC. However, ITPC is sensitive to relative and absolute trial count (Cohen, 2014a), whereas wPPC corrects for this bias because it measures similarity of phase angles between all trial pairs. The formula for wPPC as implemented in Fieldtrip function `ft_connectivity_ppc.m` is:

$$wPPC = \frac{|\sum_{j=1}^N Z|^2 - \sum_{j=1}^N |Z|^2}{|\sum_{j=1}^N |Z|^2 - \sum_{j=1}^N |Z|^2}$$

278 where Z is a complex convolution result computed at each time point using 3-cycle Morlet wavelet as
 279 described above. Conceptually, wPPC measures the extent to which the circular distance between phase
 280 angle pairs taken from different trials is non-uniformly distributed, whereas ITPC measures the extent to
 281 which a distribution of phase angles across trials at each time point is non-uniform. Analytically, wPPC is
 282 comparable to ITPC squared (Vinck et al., 2010).

283 *Statistical analyses.* Differences in the pre-stimulus and post-stimulus alpha-band power were performed
 284 taking into account individual alpha peak frequency (IAF), and its variability across parietal and occipital
 285 alpha sources (Haegens et al., 2014). Therefore, for each participant alpha band was defined as $IAF \pm 1.5$
 286 Hz separately for parietal and occipital alpha sources. Statistical comparison of pre-stimulus alpha-band
 287 power was performed on raw, i.e. non-baseline-corrected power values. Statistical comparison of pre-
 288 stimulus (-600 – 0 ms) and post-stimulus (0 – 600 ms) alpha-band time courses, and wPPC between
 289 illusion and non-illusion trials was performed using non-parametric permutation testing procedure as
 290 described further (Maris and Oostenveld, 2007).

291 First, real t-values were obtained by performing a one-sample t-test at each time point comparing
 292 the time-series of condition difference in power (or wPPC) against 0. Second, a null hypothesis
 293 distribution of t-values (permuted t-values) was created by randomly assigning condition labels
 294 (mathematically this was implemented by multiplying the condition difference time-series from a random
 295 number of participants by -1) and computing t-values. This procedure was repeated 1000 times. Third, the
 296 real t-values were z-scored using the mean and standard deviation of permuted t-values. Fourth, power (or
 297 wPPC) differences were considered statistically significant if z-scored t-value at each time point was
 298 larger than 95% of permuted t-values (i.e. z-score > 1.65). Finally, cluster-based correction was applied to
 299 correct for multiple comparisons over time points. The null hypothesis distribution of cluster sizes was
 300 obtained as follows. At each of 1000 iterations of permutation testing, the permuted t-value time series
 301 was thresholded at $p < .05$, and the maximum cluster size value (sum of t-values within a cluster) was
 302 stored. Clusters of contiguous time points in the real t-value time series were considered significant if the

size of the cluster was bigger than expected under the null hypothesis. To obtain more stable estimates from permutation testing, we ran a “meta-permutation test” by repeating pixel-level and cluster-level permutation procedure 20 times. Thus, the average of 20 real z-scored t-values, and the average of 20 cluster thresholds were used.

Statistical significance of POS was tested by comparing the observed POS value at each time-frequency point to the null hypothesis distribution of POS values, which was obtained using the following procedure. For each participant, illusion and non-illusion trials were randomly relabeled and surrogate POS was computed. This was repeated 1000 times at each time-frequency tile. Thus obtained 1000 surrogate POS values per subject were then used to compute 100,000 grand-average surrogate POS values (i.e. average across subjects). At each of 100,000 iterations, one surrogate POS value was selected for each participant and averaged across participants. Finally, p-value was computed as the proportion of grand-average surrogate POS values that exceeded empirically observed grand-average POS. This p-value indicates how empirically observed phase differences between illusion and non-illusion trials deviated from phase differences expected under null hypothesis (Fig. 5). Statistical comparison of POS frequency profile was performed by summing the observed and surrogate grand-average POS values over time (-600 to 0 ms) and computing the p-value as described above. Correction for multiple comparisons across frequencies was performed using non-parametric cluster correction procedure, equivalent to that used for alpha power and wPPC comparison between conditions.

RESULTS

Experiment 1

In Experiment 1, we empirically tested Bowen’s theoretical idea that the triple-flash illusion occurs when the delay between veridical flashes matches the period of IRF generated in response to a single flash. The two key components of Bowen’s model were determined using a psychophysical approach (subject-specific optimal delay between flashes) and EEG recordings (period of the subject-specific IRF).

327 ***Evidence for Bowen's model***

328 Replicating Bowen's results (1989), we found that illusory third flash perception depended on
 329 SOA between the two veridical flashes. The main effect of SOA on two-flash trials was significant for
 330 both coarsely ($N = 18$; $F(9, 153) = 20.86$, $p < .001$) and finely sampled SOAs ($N = 9$; $F(16, 128) = 4.85$, p
 331 $< .001$), such that perception of illusions followed an inverted u-shape function (Fig. 2A). Averaged
 332 across finely and coarsely sampled equivalent SOAs, a 75 ms delay between two veridical flashes yielded
 333 the strongest third-flash illusion: At this SOA, the illusory third flash was perceived on about half of the
 334 two-flash trials ($M = 41.41\%$, $SD = 26.83\%$). In contrast, three veridical flashes separated by the same 75
 335 ms SOA almost always were perceived as three ($M = 91.04\%$; $SD = 11.86\%$), indicating that perception
 336 of the illusory flash on two-flash trials does not result from an inability to distinguish rapidly presented
 337 stimuli. This was further supported by the main effect of SOA on three-flash trials (for finely sampled
 338 SOAs $F(1,26) = 23.27$, $p < .001$; for coarsely sampled SOAs $F(1,7) = 94.35$, $p < .001$), where perception
 339 of three flashes steadily increased as a function of SOA (Fig. 2A).

340 To directly test Bowen's theoretical account of the triple-flash illusion, we first determined each
 341 participant's impulse response function (IRF) and its period. As previously reported (VanRullen and
 342 Macdonald, 2012), white noise stimuli that have a flat frequency spectrum can be used to reveal subject-
 343 specific IRF by cross-correlating stimulus luminance time series with concurrently recorded EEG time
 344 series (Fig. 3B). As in previous reports (Ilhan and VanRullen, 2012; VanRullen and Macdonald, 2012),
 345 the cross-correlation result was oscillatory, with a period of ~ 100 ms on average and maximal amplitude
 346 over occipital electrodes (Fig. 3D). Next, we determined subject-specific SOA that maximized illusory
 347 third-flash perception by fitting symmetrical and non-symmetrical functions (Gaussian, Weibull, and
 348 exGaussian) to behavioral performance on two-flash trials. Given that exGaussian function provided the
 349 best model fits ($M(R^2_{\text{Gaussian}}) = 0.89$, $SD(R^2_{\text{Gaussian}}) = 0.09$; $M(R^2_{\text{Weibull}}) = 0.89$, $SD(R^2_{\text{Weibull}}) = 0.08$;
 350 $M(R^2_{\text{exGaussian}}) = 0.91$, $SD(R^2_{\text{exGaussian}}) = 0.08$), the peak of fitted exGaussian function was taken as a
 351 subject-specific optimal SOA (Fig. 3A).

352 The correlation between two key elements of Bowen's model – subject-specific SOA that
 353 maximized illusory perception and period of subject-specific IRF – was significant ($r_{\text{Pearson}}(25) = 0.51$,
 354 $p_{\text{two-tailed}} = .006$, $CI = [0.17 \ 0.75]$; Fig. 3C). This result is the first direct evidence for Bowen's theoretical
 355 account of the triple-flash illusion. To evaluate the influence of the outliers, we recomputed correlation
 356 for 1000 times by randomly resampling with replacement. The 95% percentile CI of these bootstrapped
 357 correlations did not include 0 ($CI_{95\%} = [0.17 \ 0.72]$), indicating robustness of the observed effect. For
 358 consistency purposes with other analyses, we also tested resistance to outliers by using a non-parametric
 359 permutation testing procedure, by randomly shuffling IRF values across participants and recomputing
 360 correlation for 1000 times (Maris and Oostenveld, 2007). The empirically observed correlation was
 361 significantly different from the null hypothesis distribution ($z\text{-score} = 2.69$, $p = .0036$), indicating
 362 robustness of the observed effect.

363 *Experiment 2*

364 In Experiment 2, we investigated the role of alpha-band oscillations in generation of the triple-flash
 365 illusion. For this we conducted an EEG experiment, for which SOA on two-flash trials was fixed at 87.5
 366 ms based on the results of Experiment 1, where maximal number of illusions on average across
 367 participants for ex-Gaussian fits was 92 ms ($SD = 17$ ms). The fixed SOA was chosen to maximize the
 368 number of illusion trials for subsequent phase-based analyses that are sensitive to the number of trials
 369 (Vinck et al., 2010; Cohen, 2014a; VanRullen, 2016a).

370 *Third-flash perception depends on individual alpha frequency*

371 Why did some participants perceive the illusion nearly half of the time and others perceived
 372 virtually none (Fig. 2B)? Based on the previously reported correlation between IRF frequency and
 373 occipital individual alpha-peak frequency (IAF; VanRullen and Macdonald, 2012), we hypothesized that
 374 variability in IAF could be related to the observed between-subject differences in proneness to perceive
 375 the third-flash illusion. Specifically, the closer the match between IAF and fixed SOA used in the EEG
 376 experiment (87.5 ms $\sim$ 11.43 Hz), the more illusions a given participant would perceive. We focused on

the frequency of task-related (as opposed to resting-state) alpha-band oscillations prior to the first flash, because IAF has been shown to be state dependent (Haegens et al., 2014). Furthermore, parietal and occipital IAFs were determined in the pre-stimulus window to avoid contamination from sensory stimulus processing that is accompanied by rapid instantaneous frequency changes in all frequency bands (Burgess, 2012). To account for IAF variability across brain regions (Klimesch, 1999; Haegens et al., 2014), we isolated occipital and parietal alpha sources using independent component analysis (for details, see Methods section). For two participants, alpha-band sources could not be determined due to small alpha peaks in the power spectrum that were indistinguishable from noise. Thus 33 participants were included in the correlation analyses. On average, there were 5.36 ICs (SD = 2.01) per subject with a clear alpha peak in the frequency spectrum and residual variance of the dipole fit (mismatch between the component scalp map and the scalp projection of a fitted dipolar source) of less than 15%. The average residual variance of the dipole fit for the selected parietal ICs was 3.77% (SD = 2.39%), and for the occipital ICs was 5.07% (SD = 3.34%).

The normalized power spectra for all participants is plotted in Fig. 4B. Average peak frequency in parietal ROI was 10.4 Hz (SD = 0.94), and in occipital ROI – 10.5 Hz (SD = 1.1). Average peak frequency of alpha oscillations across participants statistically did not differ between the two regions of interest ($t(32) = -0.625, p = .536$).

Alpha peak frequency is not stationary (Cohen, 2014b; Samaha and Postle, 2015), and fluctuations around IAF, albeit small (~ 0.04 Hz, see Samaha and Postle, 2015), can be relevant for perception. To control for the possibility that IAF estimated in the pre-stimulus window using all trials was disproportionally influenced by one class of trials (e.g. non-illusion trials), we additionally determined IAF preceding illusion and non-illusion trials separately. The subject-average (N=33) normalized power spectra for illusion and non-illusion trials separately are represented in Figure 4C, and show highly overlapping spectral profiles. A paired t-test of IAF for illusion and non-illusion trials was not significant neither for parietal ($t(32) = -0.34, p = .735$), nor occipital alpha sources ($t(32) = -.22, p =$

.827). Thus differences in IAF prior to illusion vs. non-illusion trials were negligible relative to IAF differences between parietal and occipital ROIs when compared across participants (0.6 Hz).

As hypothesized, the probability of the third-flash perception using fixed SOA was correlated with IAF: The smaller the absolute distance between parietal IAF and 11.43 Hz, the more illusions participant perceived. The correlation was significant for the parietal alpha sources ($r_{\text{Pearson}}(31) = -0.52$, $p_{\text{two-tailed}} = .002$, $\text{CI} = [-0.73 \ -0.21]$; Fig. 4C), and showed the same trend at the occipital alpha sources ($r_{\text{Pearson}}(31) = -0.33$, $p_{\text{two-tailed}} = .063$, $\text{CI} = [-0.60 \ 0.02]$; Fig. 4D). Robustness of the correlations was evaluated using bootstrapping procedure, which revealed that the 95% percentile CI of bootstrapped correlations for parietal sources did not include 0 ($\text{CI} = [-0.74 \ -0.20]$), whereas for occipital sources it did ($\text{CI} = [-0.62 \ 0.14]$). A non-parametric permutation testing procedure indicated that the empirically observed correlation was significantly different from the null hypothesis distribution ($z\text{-score} = -2.97$, $p = .0015$) for the parietal as well as occipital IAF ($z\text{-score} = -1.85$, $p = .03$).

In order to link the findings of between-subject correlations from both Experiment 1 and Experiment 2, we correlated the peak frequency of IRF determined in Experiment 1 with IAF determined in the Experiment 2. Replicating the previous findings (VanRullen and Macdonald, 2012), the peak of IAF strongly correlated with both parietal ($r_{\text{Pearson}}(16) = 0.818$, $p_{\text{two-tailed}} < .001$), and occipital IAF ($r_{\text{Pearson}}(16) = 0.774$, $p_{\text{two-tailed}} < .001$). Thus, there is a close link between oscillatory responses to a luminance increment (IRF) and endogenous posterior alpha brain rhythms (IAF).

Pre-stimulus alpha phase predicts third-flash perception

Although correlation between subject-specific optimal SOA and the period of subject-specific IRF lends support to Bowen's notion that the triple-flash illusion reflects a superposition of two oscillatory responses, it does not explain trial-to-trial variability in perception of the illusion. At the subject-specific optimal SOA, the third flash is only perceived on average half of the time (45% in the Experiment 2). To address this question, we contrasted brain activity during physically identical two-flash trials on which the third-flash was either reported, or not. All within-subject analyses were performed on

two alpha component time series (occipital and parietal), which were determined based on spectral and spatial characteristics (for details, see Methods section).

We examined the effects of pre-stimulus alpha phase on perceptual outcome by computing the phase opposition sum (POS), a measure that represents the extent to which phase distributions between two classes of trials differ (here, illusion and non-illusion trials; Busch et al., 2009; VanRullen, 2016a). Given that statistical power of POS is influenced by the absolute trial count in each class of trials (VanRullen, 2016a), and that reliability of POS can be compromised in case of unequal illusion and non-illusion trial counts (an inherent feature of all phase-based time-frequency analyses methods (Vinck et al., 2010; Cohen, 2014a), we selected only participants for which illusion and non-illusion trial counts differed by less than 10% ($N=14$). The average number of illusion trials was 333 ($SD = 43$), and non-illusion trials was 399 ($SD = 33$).

If the phase of spontaneous alpha oscillations prior to the first flash is predictive of the third-flash perception, we should observe a strong phase clustering around a certain phase angle for illusion trials accompanied by strong phase clustering around the opposite phase angle for non-illusion trials. Pre-stimulus alpha phase differed at parietal but not occipital alpha sources (Fig. 5): The POI spectrum (averaged across all time points in the pre-stimulus interval) at parietal alpha sources was statistically significant in 6-12 Hz frequency range (corrected for multiple comparisons across frequencies using cluster-based permutation testing). Although analyses were performed on alpha sources, finding phase-opposition in the alpha band is not trivial, as source-separation was not based on phase measures.

Next, we tested the effects of pre-stimulus alpha power (Fig. 6). Within-subject trial-by-trial fluctuations of pre-stimulus alpha power have been shown to affect both near-threshold stimulus perception as well as perception of illusions (Romei et al., 2008; Lange et al., 2014). Moreover, some studies reported that between-subject differences in occipital alpha power are correlated with proneness for illusory perception (Cecere et al., 2015), and performance in the perceptual discrimination tasks (Hanslmayr et al., 2007; Limbach and Corballis, 2016).

Within-subject alpha power analysis revealed significantly lower pre-stimulus alpha power on illusion vs. non-illusion trials at parietal (all p values < 0.05 in the time interval -520 – -270 ms relative to first flash, corrected for multiple comparisons using cluster-based permutation testing), but not occipital

alpha sources (all p values > 0.05 in the time interval $-600 - 0$ ms). For between-subject alpha power analyses ($N=33$), we computed subject-specific alpha power averaged across illusion and non-illusion trials in the pre-stimulus window ($-500 - -300$ ms) using two frequency bands: (1) around $IAF \pm 1.5$ Hz (for direct comparison with within-subject analyses), (2) at the stimulus presentation frequency 11.43 Hz ± 1.5 Hz. Correlations between pre-stimulus alpha power (around IAF) and overall percent of illusions was not significant, neither for parietal ($r_{\text{Pearson}}(31) = -.15$, $p_{\text{two-tailed}} = .397$), nor for occipital alpha sources ($r_{\text{Pearson}}(31) = .16$, $p_{\text{two-tailed}} = .381$). The correlation results were also non-significant when defining alpha power around the stimulation frequency (i.e. 11.43 Hz ± 1.5 Hz): The correlation at parietal alpha sources was $r(31) = -.028$ ($p_{\text{two-tailed}} = .879$), and occipital alpha sources was $r(31) = .195$ ($p_{\text{two-tailed}} = .276$).

To rule out that IAF estimation was affected by between-subject differences in alpha power, we also correlated IAF with absolute pre-stimulus alpha power. None of the correlations were significant: The correlation with parietal alpha power around IAF was $r_{\text{Pearson}}(31) = 0.01$ ($p_{\text{two-tailed}} = .945$), and around 11.43 was $r_{\text{Pearson}}(31) = .20$ ($p_{\text{two-tailed}} = .266$); the correlation with occipital alpha power around IAF was $r_{\text{Pearson}}(31) = 0.21$ ($p_{\text{two-tailed}} = .243$), and around 11.43 was $r_{\text{Pearson}}(31) = .27$ ($p_{\text{two-tailed}} = .127$). Together, these results demonstrate that trial-by-trial alpha power fluctuations (rather than between-subject differences in alpha power) at parietal but not occipital alpha sources are one of the elements determining the probability to perceive an illusion on a given trial.

Illusory perception is associated with stronger post-stimulus local phase alignment

According to Bowen's theoretical model, presentation of the second flash in-phase with the oscillatory IRF evoked by the first flash would result in perfect superposition of the two IRFs and hence the perception of an illusory third flash. Whenever, for any reason (e.g. non-optimal SOA, non-optimal alpha phase at the first-flash onset, variability in stimulus-evoked oscillatory alpha phase, etc.), the second flash does not occur in-phase with the oscillatory response evoked by the first flash, then the model stipulates that third-flash perception would be less likely. Thus, we predicted more precise phase alignment on illusion than non-illusion trials in the post-stimulus window. We assessed phase alignment by computing weighted pair-wise phase consistency metric (wPPC; Vinck et al., 2010) at the frequency of the two veridical flashes (11.43 Hz ~ 87.5 ms). As illustrated in Figure 7, we found significantly higher wPPC for illusion than for non-illusion trials at parietal but not occipital alpha sources, indicating higher phase consistency. We also compared post-stimulus alpha power to rule out the possibility that wPPC

485 differences were a result of less accurate phase estimation due to low alpha power (Cohen, 2014a). We
 486 observed a typical decrease in alpha power related to stimulus processing, but this effect did not differ
 487 between illusion and non-illusion trials.

488

489 **DISCUSSION**

490 We validated and extended the original theoretical account of the triple-flash illusion proposed by
 491 Bowen (1989), according to which the illusory third-flash percept arises when the delay between the two
 492 veridical flashes matches the period of a hypothetical oscillatory impulse response function (IRF)
 493 generated in response to each stimulus. In Experiment 1, we demonstrated that the subject-specific inter-
 494 flash delay for which the illusory perception is maximized was strongly correlated with the period of the
 495 oscillatory IRF, which reverberates at ~ 10 Hz (Fig. 3C). The subject-specific IRF was derived from EEG
 496 responses to white-noise luminance sequences by cross-correlating the two signals (VanRullen and
 497 Macdonald, 2012). In Experiment 2, when fixing the inter-flash delay (87.5 ms) for all participants, we
 498 demonstrated that individual alpha peak frequency (IAF) of parietal as compared to occipital alpha
 499 sources was more strongly correlated with the overall proportion of illusory percepts: The closer
 500 participant's parietal alpha peak was to 11.43 Hz (the 87.5 ms delay, expressed in Hz), the more illusions
 501 were perceived. Together, these results point to an active or 'driving' (as opposed to modulatory) role of
 502 alpha-band oscillations in perception, and reveal that alpha-band reverberations to a single stimulus have
 503 direct consequences on perception spanning several subsequent alpha cycles. Moreover, these findings
 504 emphasize the importance of using the "individual differences" approach when studying perceptual cycles
 505 and their associated oscillatory signatures.

506 Parietal and occipital alpha sources were estimated from ICA and single dipole fitting – a
 507 combination of methods that appropriately dissociates highly spatially adjacent oscillatory sources
 508 (Tollner et al., 2017). Although ICs are often dipolar (Delorme et al., 2012), the anatomical dissociation
 509 of occipital vs. parietal sources should be interpreted cautiously, considering that dipole localization was
 510 based on 64-channel EEG using standard electrode locations, and a standard anatomical head model

(although standard head models can provide reasonably high localization accuracy; Fuchs et al., 2002). However, a differential role of occipital vs. parietal alpha in perception has been reported previously and is consistent with our findings. In a discrimination task, for example, lower alpha power in parietal sources (BA 7) preceded correct trials, and was interpreted to reflect information gating from occipital to dorsal parietal areas controlled by top-down effects of attention (van Dijk et al., 2008). Importantly, this parietal alpha source was distinct from the occipital alpha source identified from the resting-state recordings. Relatedly, behavioral effects of 10 Hz rTMS in another visual discrimination task were observed only when stimulating parietal, but not occipital areas (Jaegle and Ro, 2014). Although this effect might be due to stronger entrainment effects in parietal compared to occipital rTMS, the proposed active role of parietal alpha in modulating visual representations in lower visual areas could account for the differential rTMS effects (Foxe and Snyder, 2011; Palva and Palva, 2011; Kwon et al., 2016). Pre-stimulus alpha phase in fronto-parietal areas has also been demonstrated to affect the connectivity between occipital and parietal areas, with certain pre-stimulus alpha phases associated with increased connectivity and better near-threshold stimulus detection (Hanslmayr et al., 2013).

Only two empirical studies of the triple-flash illusion, to our knowledge, have been conducted since Bowen's original report; they compared the optimal SOA of the triple-flash illusion between healthy controls and schizophrenia patients (Norton et al., 2008; Chen et al., 2014). In both studies, the average SOA that maximized illusory perception was longer in the schizophrenia patient group (130-150 ms) than in the control group (90-110 ms). Following Bowen's model, the authors speculated that such differences could result from temporal dilation of the IRF in schizophrenia patients. Considering both our between-subject correlation analyses results (Fig. 3-4), and previous reports of slower IAF in patients with chronic schizophrenia and schizophrenia symptoms (Canive et al., 1998; Harris et al., 2006), a slowing down of occipito-parietal alpha rhythms seems a conceivable explanation for the Norton et al. and Chen et al. findings.

Although Bowen's theoretical model explains the mechanics behind the triple-flash illusion (Fig. 1B), it does not explain why at the subject-specific optimal SOA, an illusory third-flash is only perceived

half of the time (45% on average). We hypothesized that this probabilistic nature of the illusion could be related to moment-to-moment fluctuations in occipito-parietal alpha phase and power that are known to be perceptually relevant (VanRullen et al., 2011; Kleinschmidt et al., 2012). In the EEG experiment (Experiment 2), we found that trial-to-trial variability in perception of the illusion was indeed related to the pre-stimulus alpha phase at parietal but not occipital alpha sources, such that illusion and non-illusion trials were associated with opposite alpha phases (Fig. 5). Illusion trials were also preceded by significantly lower pre-stimulus alpha-band power at parietal alpha sources (Fig. 6). These findings are in accordance with previous reports linking relatively lower occipito-parietal alpha power and certain alpha phases to higher cortical excitability (Romei et al., 2008; Mathewson et al., 2009; Dugue et al., 2011; Lange et al., 2013). Variations in occipito-parietal alpha power are related not only to veridical but also illusory perception: Illusory percepts are more frequent when occipito-parietal alpha power is low, both within- and between-subjects (VanRullen et al., 2006; Lange et al., 2014; Cecere et al., 2015).

Presentation of stimuli in phase with endogenous alpha oscillations results in stronger phase consistency across trials as compared to jittered stimulation (Thut et al., 2011; Spaak et al., 2014; Notbohm et al., 2016). In the post-stimulus period, we found higher phase consistency (as measured by wPPC) for illusion than non-illusion trials at 11.43 Hz (the two-flash trial SOA of 87.5 ms, translated into Hz) for parietal but not occipital alpha sources (Fig. 7). The wPPC differences between illusion and non-illusion trials, as well as pre-stimulus phase-opposition effects, appear complementary with Bowen's theoretical model, which posits that illusory third-flash perception is associated with an enhancement of response amplitude resulting from phase-aligned oscillatory responses to each stimulus. Specifically, when the first stimulus appears in phase with ongoing pre-stimulus alpha oscillations (i.e., at the "good" pre-stimulus phase), there is no or relatively little phase re-alignment (Fellinger et al., 2011; Gruber et al., 2014); the second stimulus thus also arrives in phase with the ongoing oscillations and in phase with the oscillatory response generated to the first stimulus, resulting in high response amplitude and a third-flash percept. However, when the first stimulus arrives slightly or completely out of phase with the ongoing

562 alpha oscillations (i.e., at the “bad” pre-stimulus phase), the resulting phase alignment in response to both
 563 stimuli is less precise (weaker wPPC), and thus only two flashes are perceived.

564 We demonstrate that perception (in a broad sense) is not only influenced by spontaneous pre-
 565 stimulus alpha oscillations (Busch et al., 2009; Iemi et al., 2017), but that oscillatory events related to
 566 stimulus processing on one cycle have downstream perceptual effects on multiple subsequent alpha
 567 cycles. Importantly, these non-linear perceptual effects are most pronounced when the lag between two
 568 successive stimuli matches the rhythm of task-related alpha-band oscillations. This is akin to the process
 569 of entrainment of brain oscillations to rhythmic external stimuli (Thut et al., 2011), which is most
 570 effective when periodic light flashes are “in-sync” with individual alpha peak frequency (Adrian and
 571 Matthews, 1934; Notbohm et al., 2016).

572 Current theoretical frameworks on the functional role of alpha oscillations posit that alpha
 573 oscillations influence sensory and cognitive processes in a pulsed manner (Klimesch et al., 2007; Jensen
 574 and Mazaheri, 2010; Mathewson et al., 2011), with strong alpha oscillations reflecting physiological
 575 inhibition which is more pronounced at certain phases of the alpha cycle (Sadaghiani and Kleinschmidt,
 576 2016). In the context of perceptual tasks (detection and discrimination), these theories emphasize the
 577 modulatory role of the alpha rhythm in signal processing. Consistent with the idea that the alpha rhythm
 578 represents “pulsed inhibition” (Mathewson et al., 2011), we found that illusory perception was associated
 579 with low alpha power (Fig. 6), and with a specific alpha phase (Fig. 5). However, our findings of non-
 580 linear perceptual effects (triple-flash illusion) which result from the interaction between the subject-
 581 specific endogenous alpha rhythm and the stimulus rhythm cannot be explained when considering alpha
 582 oscillations only as a modulatory rhythm. Instead, these results point to the driving power of alpha-band
 583 oscillations – perception of illusory stimuli without corresponding sensory input in the same or different
 584 modality – which had not been experimentally demonstrated before.

585 In conclusion, using the triple-flash illusion – a third illusory flash perception when only two
 586 veridical ones are presented, separated by ~100 ms – we demonstrate that alpha-band oscillations not only
 587 modulate perception but have a driving impact, which can make one perceive something that is not there.

Figure Legends

Figure 1. Trial structure and Bowen's theoretical model. (A) On each trial two or three high-contrast circle stimuli were presented in rapid succession above the fixation dot. Stimulus onset asynchrony (SOA) on two-flash trials was either variable (Experiment 1), or fixed (Experiment 2); three-flash trial SOA was variable in both experiments. (B) Oscillatory model of brain response to two consecutive light flashes (adapted from Bowen 1989). Thin black and blue lines represent an impulse response function (IRF) to the first and the second flash respectively; the thick green line represents a linear sum of IRFs to two flashes. The horizontal dashed lines represent a hypothetical perceptual threshold which, when exceeded, gives rise to the perception of a flash. In the top and the bottom panels the delay between veridical flashes is suboptimal for the third-flash perception to occur (i.e. too short, SOA = 50 ms, and too long SOA = 150 ms). The middle panel represents a scenario when the second flash is presented ~100 ms relative to the first flash, resulting in a perfect superposition of the two IRFs and a third-flash percept.

Figure 2. Behavioral performance. Percentage of three-flash percepts for two- and three-flash trials as a function of stimulus onset asynchrony (SOA) between veridical flashes. (A) Behavioral results of psychophysics Experiment 1, where fine (17 SOAs for two- and 27 SOAs for three-flash trials) and coarse (10 SOAs for the two- and 8 SOAs for the three-flash trials) were used. Dotted arrows indicate two-flash trial SOA that produced on average the maximum number of illusions. (B) Behavioral results of the EEG experiment (Experiment 2), where the two-flash trial SOA was fixed at 87.5 ms. Grey circles represent average percent of illusions for each participant (N=35). Error bars show standard error of the mean.

Figure 3. Relationship between period of impulse response function (IRF) and optimal delay between flashes. (A) Behavioral data from two participants (S1 and S2) showing the probability of third-flash perception on two-flash trials as a function of SOA (black lines), and exGaussian fits used to determine subject-specific optimal SOA (green and orange lines). (B) Single-subject IRFs at Oz electrode (same participants as in panel A). (C) Correlation (across participants, N=27) between period of IRF at Oz electrode and optimal SOA. Dashed curves represent 95% confidence intervals around the slope of regression line. Green and orange dots indicate participants' data depicted in panels A and B. (D) Subject-average power spectrum of IRF at Oz and topographical map of 10 ± 3 Hz power indicating that IRF is strongest around Oz (white circle). Subject-average power spectrum of IRF at Oz electrode with a peak centered at 10 Hz. Light red areas represent standard error of the mean.

Figure 4. Relationship between the frequency of alpha oscillations and third-flash illusory percepts. (A) Dipole locations for anatomical reference coordinates for parietal and occipital ROIs (left-most panel). Locations of equivalent dipoles for parietal and occipital alpha independent components (ICs). (B)

627 Power spectra of alpha parietal and occipital IC time series in the pre-stimulus window (-1000 – 0 ms,
 628 where 0 is the first flash), normalized to the power of each participant's alpha peak (for comparability
 629 across participants). Each row represents a participant, color corresponds to normalized spectral power,
 630 and the white dot denotes the individual's alpha peak frequency. The power spectra of exemplar subjects
 631 S1 and S2 depicted in Figure 3A, B are highlighted with dashed rectangles. Inset topographical maps
 632 represent scalp projections of parietal and occipital equivalent dipole centroids, revealing more central
 633 and anterior projection for parietal ICs. (C) Subject-average normalized power spectra of occipital and
 634 parietal IC time series in the pre-stimulus window, illustrating that IAF defined from all vs. separate
 635 classes of trials did not differ. (D, E) Correlations (across participants, N=33) between individual alpha
 636 peak frequency (IAF) and overall percent of illusions, indicating that the proportion of illusions depended
 637 more strongly on the match between inter-flash delay (11.43 Hz ~ 87.5 ms) and IAF at parietal (D) than
 638 occipital alpha sources (E).

639

640 **Figure 5. Pre-stimulus alpha phase differences between illusion and non-illusion trials.** (A) Time-
 641 frequency representation of p-values (for parietal alpha sources) computed as a proportion of surrogate
 642 phase opposition values (distribution of phase opposition values expected under null hypothesis) that
 643 exceeded empirically observed phase opposition values. (B) Frequency profile of phase opposition
 644 (averaged over -600 – 0 ms pre-stimulus time window). Grey shaded area represents frequencies at which
 645 the observed phase opposition frequency profile was significantly different from the frequency profile of
 646 surrogate phase opposition values (corrected for multiple comparisons across frequencies using cluster-
 647 based permutation testing). The p-value of 0.05 is marked by a horizontal line.

648

649 **Figure 6. Pre-stimulus alpha power differences between illusion and non-illusion trials.** (A) Time-
 650 frequency representation of alpha power differences at parietal sources. (B) Time-courses of alpha power
 651 (IAF $\pm$ 1.5 Hz) for illusion (dashed lines) and non-illusion (solid lines) trials at occipital and parietal alpha
 652 sources, demonstrating that alpha power differences were present at parietal but not occipital alpha
 653 sources. Grey shaded area represents the time interval where statistically significant differences between
 654 the two trial groups was observed (corrected for multiple comparisons using cluster-based permutation
 655 testing).

656

657 **Figure 7. Pairwise phase consistency for illusion and non-illusion trials.** Time-courses of weighted
 658 pair-wise phase consistency (wPPC) at 11.43 Hz for illusion and non-illusion trials plotted separately for
 659 parietal (left panel) and occipital (right panel) alpha sources. Black bar on the x-axis represents time
 660 points at which wPPC for illusion vs. non-illusion trials was significantly different (corrected for multiple
 661 comparisons using cluster-based permutation testing).

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